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Latent variable nonparametric cointegrating regression. (English summary)

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This article investigates statistical properties of a nonparametric cointegrating regression model with a latent regressor. Let  $(y_t, z_t)$  be two nonstationary variables that are linked according to the nonparametric regression model

$$y_t = g(z_t) + \eta_t, \quad t = 1, \dots, n,$$

where  $g(\cdot)$  is an unknown function and  $\eta_t$  is a zero-mean stationary disturbance that may be correlated with  $z_t$ . Throughout it is assumed that the true regressor  $z_t$  is unobservable, whereas its proxy  $x_t$  is available. Such circumstances arise when  $z_t$  is measured with error or given in filtered forms as in the mixed-data-sampling (MIDAS) regression or Hodrick-Prescott (HP) filtering.

In this scenario, it is a common practice to use  $x_t$  in place of  $z_t$  to estimate  $g(x)$  nonparametrically as

$$\hat{g}(x) = \frac{\sum_{t=1}^n y_t K\{(x_t - x)/h\}}{\sum_{t=1}^n K\{(x_t - x)/h\}},$$

where  $K(\cdot)$  is a nonnegative kernel function and  $h$  is a bandwidth. To explore the limit theory of  $\hat{g}(x)$ , observe that

$$(1) \quad \hat{g}(x) = \frac{\sum_{t=1}^n \eta_t K\{(x_t - x)/h\}}{\sum_{t=1}^n K\{(x_t - x)/h\}} + \frac{\sum_{t=1}^n g(z_t) K\{(x_t - x)/h\}}{\sum_{t=1}^n K\{(x_t - x)/h\}}.$$

While the first term can be handled in a manner analogous to correctly specified cases, the second term involves  $S_n := \sum_{t=1}^n g(z_t) K\{(x_t - x)/h\}$ , the limit theory of which depends on the linkage between two nonstationary variables,  $(x_t, z_t)$ , and  $n$ , and is new in the literature.

To obtain a general result on the limit behavior of  $S_n$ , let  $a_t$  be asymptotically cointegrated with  $x_t$  so that

$$a_t = \gamma_{nt} x_t + w_{nt} + u_t,$$

where  $\gamma_{nt}$  is a sample-size-dependent coefficient with a limit  $\gamma_0$  (e.g., an estimated coefficient),  $w_{nt}$  is an asymptotically negligible misspecification error, and  $u_t$  is a stationary disturbance. Also let  $\xi_t$  be the innovation for  $x_t$  and  $d_n := \sqrt{\text{Var}(\sum_{t=1}^n \xi_t^2)}$ . Then, it is demonstrated that for a class of continuous functions  $f(\cdot)$ , a given  $x$  and the bandwidth satisfying  $h + d_n/(nh) \rightarrow 0$  as  $n \rightarrow \infty$ ,

$$(2) \quad \frac{d_n}{nh} \sum_{t=1}^n f(a_t) K\left(\frac{x_t - x}{h}\right) \xrightarrow{d} E\{f(\gamma_0 x + u_1)\} L_Z(1, 0),$$

where  $L_Z(t, x)$  is the continuous local time for a fractional Ornstein-Uhlenbeck process  $Z_t$ .

The limit behavior of  $\hat{g}(x)$  can be obtained by applying (2) to the second term of (1). Suppose that  $(x_t, z_t)$  has the linkage

$$x_t = \alpha_{nt} z_t + w_{nt} + u_t,$$

where  $\max_{1 \leq k \leq n} |\alpha_{nk}^{-1} - 1| \rightarrow 0$  as  $n \rightarrow \infty$ , and  $x_t$  (and thus  $z_t$ ) is correlated with  $\eta_t$ . Then, the first term of (1) is shown to be  $O_p\{(nh^2)^{1/4}\}$ , whereas reading  $f$  as  $g$  in (2) yields

$$\left\{ \frac{d_n}{nh} \sum_{t=1}^n K\left(\frac{x_t - x}{h}\right), \frac{d_n}{nh} \sum_{t=1}^n g(z_t) K\left(\frac{x_t - x}{h}\right) \right\} \xrightarrow{d} \{L_Z(1, 0), E\{g(x - u_1)\}L_Z(1, 0)\}.$$

Therefore, applying the continuous mapping theorem to the second term of (1), we finally have

$$\widehat{g}(x) = O_p\{(nh^2)^{1/4}\} + E\{g(x - u_1)\} + o_p(1) \xrightarrow{p} E\{g(x - u_1)\} =: g_1(x).$$

The function  $g_1(x)$  is a weighted average of  $g$  locally around  $x$  with weights determined by the stationary distribution of the measurement error  $u$ . Also observe that  $\widehat{g}(x)$  is usually inconsistent because  $g(x) \neq g_1(x)$  in general. However, when  $g$  is linear,  $g(x) = g_1(x)$  holds and thus  $\widehat{g}(x)$  becomes consistent. This corresponds to the case in which the true regressor  $z_t$  is measured with a stationary error in a linear cointegrating regression.

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